

Mobile Communications

TCS 455

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Lecture 12

Office Hours:

BKD 3601-7

Tuesday 14:00-16:00

Thursday 9:30-11:30

Announcements

- Read
 - Chapter 3: 3.1 – 3.2, 3.5.1, 3.6, 3.7.2
 - Posted on the web
 - Appendix A.1 (Erlang B)
- Due date for HW3: Dec 18

Big Picture

S = total # available duplex radio channels for the system



Frequency reuse with **cluster size N**

“Capacity”

$$C = \frac{A_{\text{total}}}{A_{\text{cell}}} \times \frac{S}{N}$$

Tradeoff

$$\frac{S}{I} \approx \frac{kR^{-\gamma}}{K \times (kD^{-\gamma})} = \frac{1}{K} \left(\frac{D}{R} \right)^\gamma = \frac{1}{K} (\sqrt{3N})^\gamma$$

m = # channels allocated to each cell.

- Omni-directional: $K = 6$
- 120° Sectoring: $K = 2$
- 60° Sectoring: $K = 1$



Trunking

$$P_b = \frac{\frac{A^m}{m!}}{\sum_{i=0}^m \frac{A^i}{i!}}$$

λ = Average # call attempts/requests per unit time

A = **traffic intensity** or load [Erlangs] = $\frac{\lambda}{\mu}$

$\frac{1}{\mu} = H$ = Average call length

Erlang-B formula

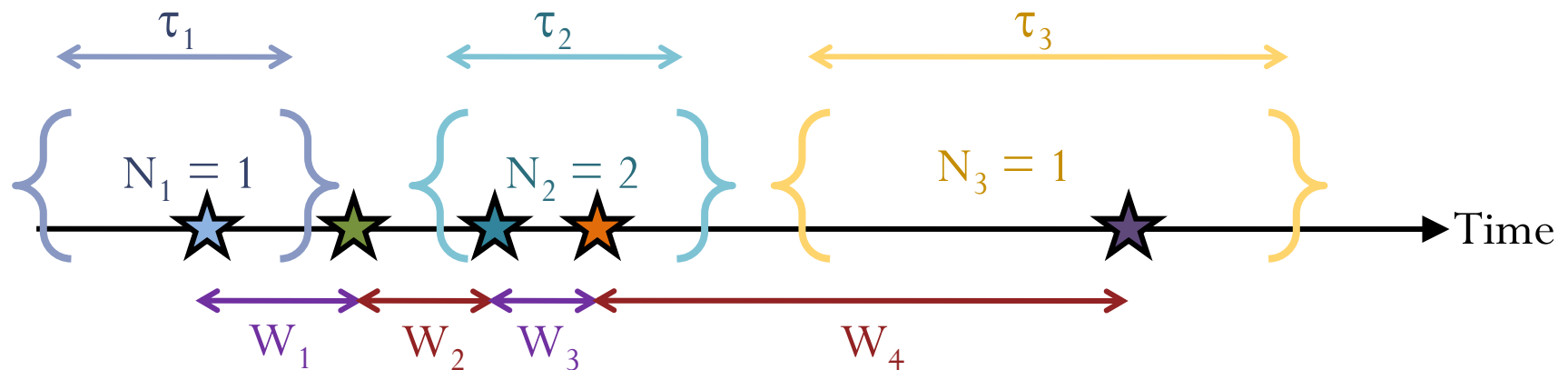
Call blocking probability

Assumption

- **Blocked calls cleared**
 - Offers no queuing for call requests.
 - For every user who requests service, it is assumed there is no setup time and the user is given immediate access to a channel if one is available.
 - If no channels are available, the requesting user is blocked without access and is free to try again later.
- **Calls arrive as determined by a *Poisson process*.**
- There are memoryless arrivals of requests, implying that all users, including blocked users, may request a channel at any time.
- There are an infinite number of users (with finite overall request rate).
 - The finite user results always predict a smaller likelihood of blocking. So, assuming infinite number of users provides a conservative estimate.
- **The duration of the time that a user occupies a channel is exponentially distributed**, so that longer calls are less likely to occur.
- There are m channels available in the trunking pool.
 - For us, $m =$ the number of channels for a cell (C) or for a sector

Poisson Process

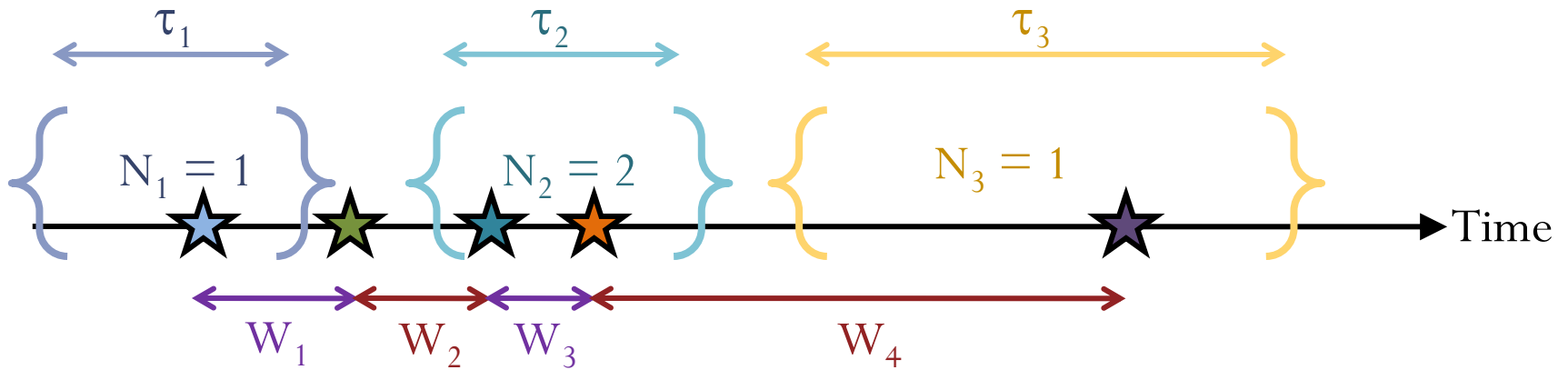
The number of arrivals N_1 , N_2 and N_3 during non-overlapping time intervals are independent **Poisson** random variables with mean $= \lambda \times$ the length of the corresponding interval.



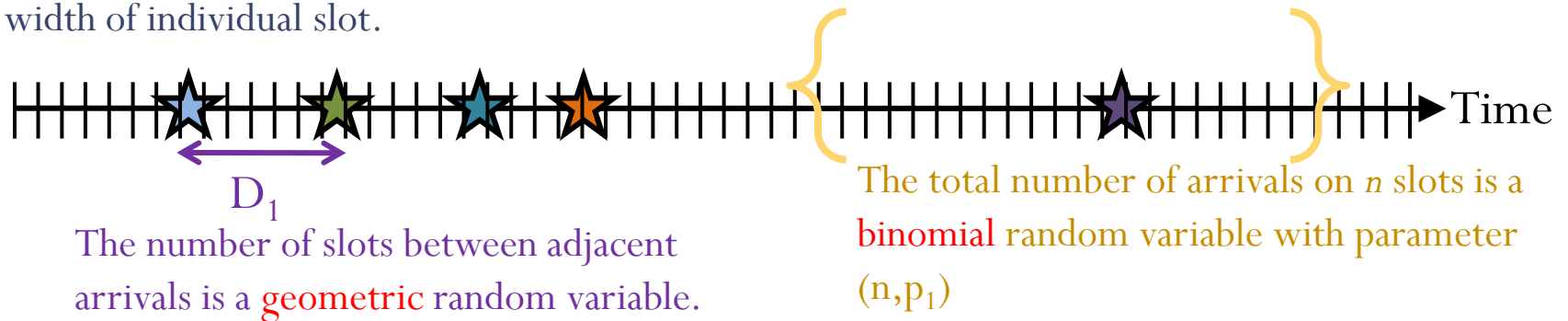
The lengths of time between adjacent arrivals $W_1, W_2, W_3 \dots$ are i.i.d. **exponential** random variables with mean $1/\lambda$.

Small Slot Analysis (Poisson Process)

- Aka discrete time approximation



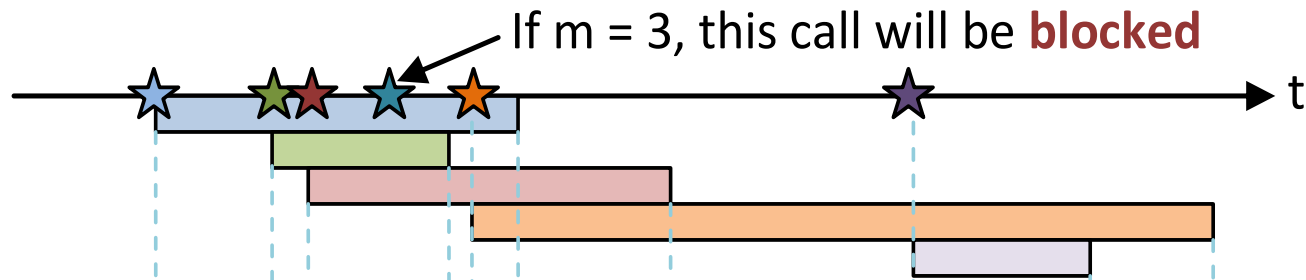
In the limit, there is at most one arrival in any slot. The numbers of arrivals on the slots are i.i.d. **Bernoulli** random variables with probability p_1 of exactly one arrivals = $\lambda\delta$ where δ is the width of individual slot.



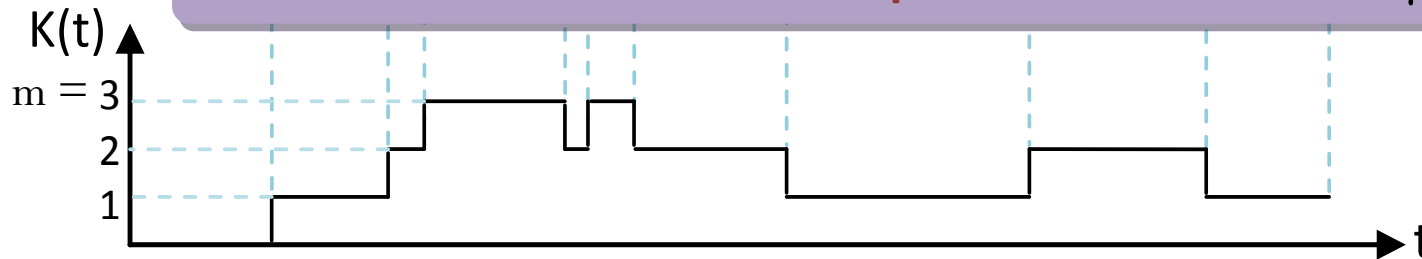
In the limit, as the slot length gets smaller, geometric \longrightarrow exponential
 binomial \longrightarrow Poisson

Assumption (2)

The call request process is **Poisson** with rate λ



The duration of calls are i.i.d. **exponential** r.v. with rate μ .

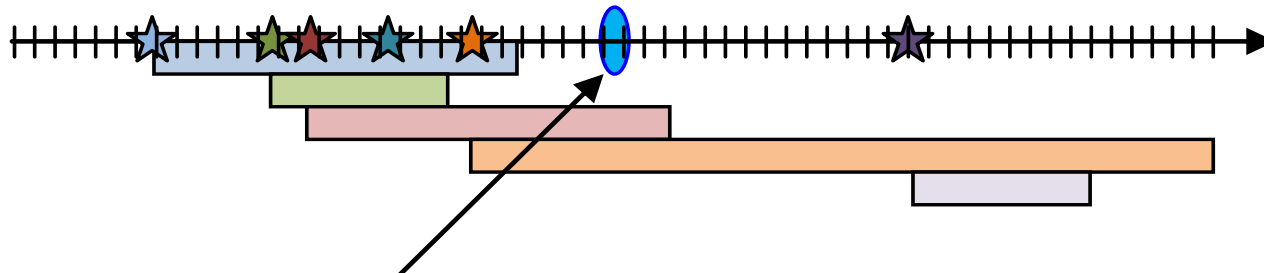


$K(t)$ = "state" of the system
= the number of used channel at time t

We want to find out what proportion of time the system has $K = m$.

Small Slot Analysis (Erlang B)

Suppose each slot duration is δ .



- Consider the i^{th} small slot.
- Let $K_i = k$ be the value of K at the beginning of this time slot.
- $k = 2$ in the above figure.
- Then, K_{i+1} is the value of K at the end of this slot which is the same as the value of K at the beginning of the next slot.

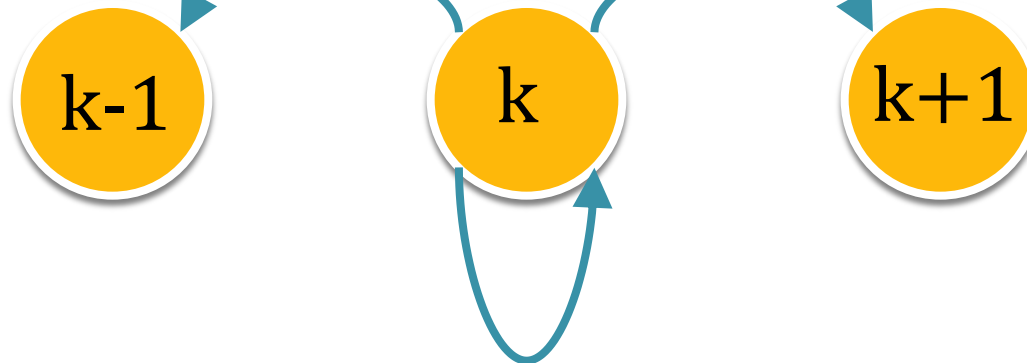
- $P[0 \text{ new call request}] \approx 1 - \lambda\delta$
- $P[1 \text{ new call request}] \approx \lambda\delta$
- $P[0 \text{ old-call end}] \approx (1 - \mu\delta)^k \approx 1 - k\mu\delta$
- $P[1 \text{ old-call end}] \approx k\mu\delta(1 - \mu\delta)^{k-1} \approx k\mu\delta$

How do these events affect K_{i+1} ?

Small slot Analysis (2)

$$K_{i+1} = K_i + (\# \text{ new call request}) - (\# \text{ old-call end})$$

$$(1 - \lambda\delta)(k\mu\delta) \approx k\mu\delta \quad (\lambda\delta)(1 - k\mu\delta) \approx \lambda\delta$$



$$(1 - \lambda\delta)(1 - k\mu\delta) + (\lambda\delta)(k\mu\delta) \approx 1 - \lambda\delta - k\mu\delta$$

The labels on the arrows are probabilities.

$$P[0 \text{ new call request}] \approx 1 - \lambda\delta$$

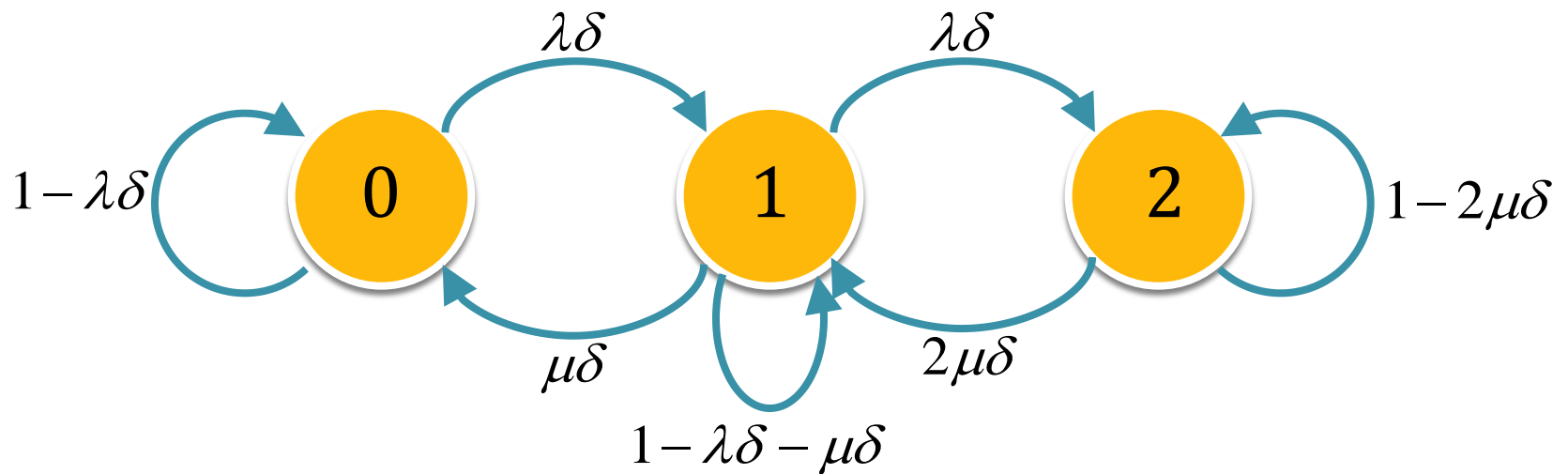
$$P[1 \text{ new call request}] \approx \lambda\delta$$

$$P[0 \text{ old-call end}] \approx 1 - k\mu\delta$$

$$P[1 \text{ old-call end}] \approx k\mu\delta$$

Small slot Analysis: Markov Chain

- Case: $m = 2$



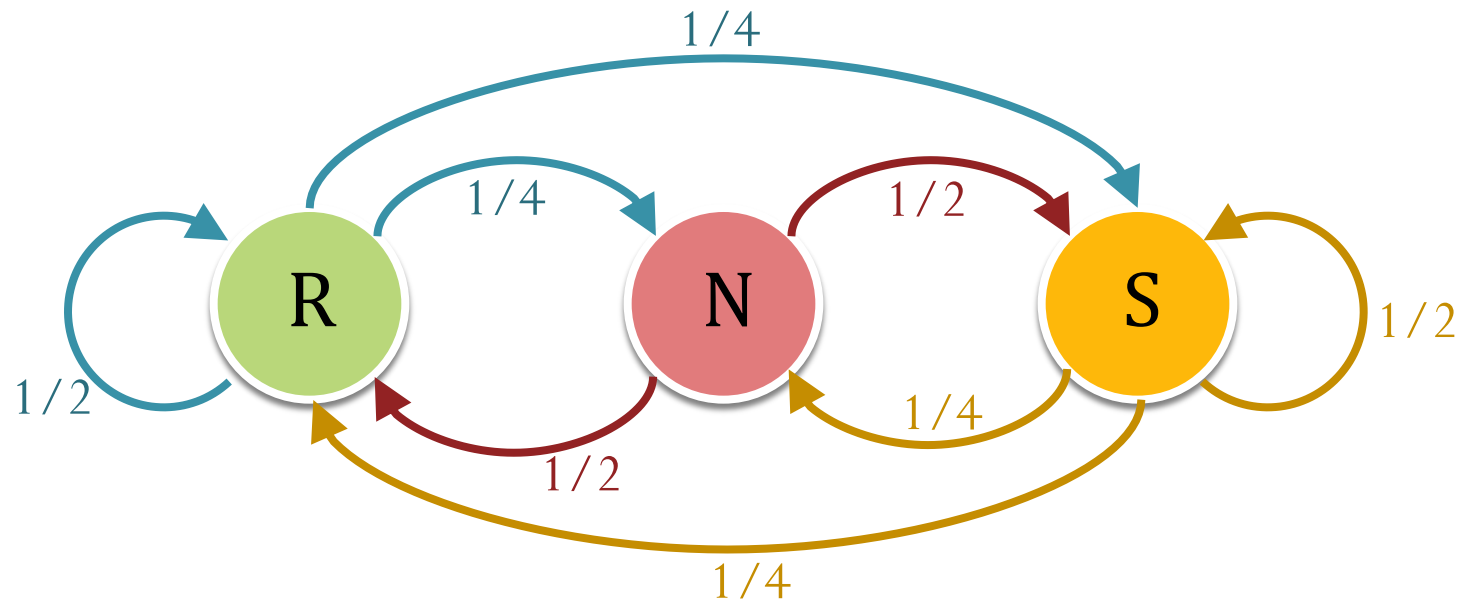
Markov Chain

- Markov chains model many phenomena of interest.
- We will see one important property: **Memoryless**
 - It retains no memory of where it has been in the past.
 - Only the current state of the process can influence where it goes next.
- Very similar to the *state transition diagram* in digital circuits.
 - In digital circuit, the labels on the arrows indicate the input/control signal.
 - Here, the labels on the arrows indicate transition probabilities. (If the system is currently at a particular state, where would it go next on the next time slot?)
- We will focus on **discrete time Markov chain**.

Example: The Land of Oz

- Land of Oz is blessed by many things, but not by good weather.
 - They never have two nice days in a row.
 - If they have a nice day, they are just as likely to have snow as rain the next day.
 - If they have snow or rain, they have an even chance of having the same the next day.
 - If there is change from snow or rain, only half of the time is this a change to a nice day.
- If you visit the land of Oz next year for one day, what is the chance that it will be a nice day?

State Transition Diagram



R = Rain
N = Nice
S = Snow

Markov Chain (2)

- Let K_i be the weather status for the i^{th} day (from today).
- Suppose we know that it is snowing in the land of Oz today. Then

$$K_0 = S$$

where S means snow.

- **Goal:** We want to know whether $K_{365} = N$ where N means nice.
- Of course, the weather are controlled probabilistically; so we can only find **$P[K_{365} = N]$** .
- From the specification (or from the state transition diagram), we know

that

$$P[K_1 = R] = \frac{1}{4}, \quad P[K_1 = N] = \frac{1}{4}, \quad P[K_1 = S] = \frac{1}{2}$$

- Define vector

$$\vec{p}(i) = [P[K_i = R] \quad P[K_i = N] \quad P[K_i = S]]$$

- Then,

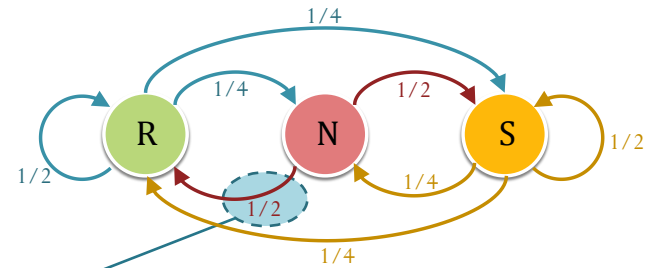
$$\vec{p}(0) = [0 \quad 0 \quad 1] \text{ and } \vec{p}(1) = \begin{bmatrix} \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{bmatrix}$$

The Land of Oz: Transition Matrix

$$\bar{p}(i+1) = \bar{p}(i) \times P$$



$$P = \begin{matrix} & \begin{matrix} R & N & S \end{matrix} \\ \begin{matrix} R \\ N \\ S \end{matrix} & \begin{bmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{bmatrix} \end{matrix}$$



$$P[K_{i+1} = \text{rain} | K_i = \text{normal}]$$

$$\bar{p}(n) = \bar{p}(0) \times P^n$$

$$\bar{p}(2) = [0.3750 \quad 0.1875 \quad 0.4375]$$

$$\bar{p}(3) = [0.3906 \quad 0.2031 \quad 0.4063]$$

$$\bar{p}(5) = [0.3994 \quad 0.2002 \quad 0.4004]$$

$$\bar{p}(7) = [0.4000 \quad 0.2000 \quad 0.4000] = \bar{p}(8) = \bar{p}(9) = \bar{p}(10) = \dots = \bar{p}(365)$$

Finding P^n for “large” n

$$P = \begin{bmatrix} 1 & 1 & 1 \\ 2 & 4 & 4 \\ 1 & 0 & 1 \\ 2 & & 2 \\ 1 & 1 & 1 \\ 4 & 4 & 2 \end{bmatrix} \quad \rightarrow \quad P^2 = \begin{bmatrix} 0.4375 & 0.1875 & 0.3750 \\ 0.3750 & 0.2500 & 0.3750 \\ 0.3750 & 0.1875 & 0.4375 \end{bmatrix}$$
$$P^3 = \begin{bmatrix} 0.4063 & 0.2031 & 0.3906 \\ 0.4063 & 0.1875 & 0.4063 \\ 0.3906 & 0.2031 & 0.4063 \end{bmatrix}$$
$$P^5 = \begin{bmatrix} 0.4004 & 0.2002 & 0.3994 \\ 0.4004 & 0.1992 & 0.4004 \\ 0.3994 & 0.2002 & 0.4004 \end{bmatrix}$$
$$P^7 = \begin{bmatrix} 0.4000 & 0.2000 & 0.4000 \\ 0.4000 & 0.2000 & 0.4000 \\ 0.4000 & 0.2000 & 0.4000 \end{bmatrix} = P^8 = P^9 = P^{10} = \dots$$

Land of Oz: Answer

- Recall that

$$\bar{p}(n) = \bar{p}(0) \times P^n$$

- So,

$$\bar{p}(7) = \bar{p}(0)P^7 = [0.4 \quad 0.2 \quad 0.4]$$

- Note that the above result is true regardless of the initial $\bar{p}(0)$

- $\bar{p}(365) = \bar{p}(0)P^{365} = [0.4 \quad 0.2 \quad 0.4]$

$\mathbf{P}[K_{365} = \mathbf{N}]$



Tip: Alternative look

$$P = \begin{bmatrix} 2/5 & 3/5 \\ 1/2 & 1/2 \end{bmatrix}$$

